Quantitative Predictors and Interactions

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So far, we've focused on only quantitative predictors.

Often, data sets have one or more qualitative predictors.

We need to consider how to fit these into a numeric model fitting context.

Qualitative Predictors with Two Levels

Consider the variable Own from the Credit data.

```
credit <- read.csv("C:/Users/cappiello/OneDrive - California State Univers
own <- as.factor(credit$Own)
summary(own)
```
No Yes ## 193 207

To put this into a regression model, we use a dummy variable:

 $x_i = I$ (the *i*th person owns a house)

Qualitative Predictors with Two Levels

This results in the model

$$
y_i = \beta_0 + \beta_1 x_i + \epsilon_i
$$

which takes values

$$
\triangleright \ \beta_0 + \beta_1 + \epsilon_i \text{ if the } i\text{th person owns a house.}
$$

and

 \triangleright $\beta_0 + \epsilon_i$ if the *i*th person does not own a house.

So β_1 is the average difference in credit card balance between owners and non-owners.

Qualitative Predictors with Two Levels

```
summary(lm(Limit ~ Own, data = credit))
```


Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2311 on 398 degrees of freedom Multiple R-squared: 8.83e-05, Adjusted R-squared: -0.002424 F-statistic: 0.03515 on 1 and 398 DF, p-value: 0.8514

Qualitative Predictors with More than Two Levels

Consider the variable region from the Credit data.

East South West ## 99 199 102

We can represent this by constructing two dummy variables.

 $x_{i,1} = I$ (*i*th person is from the South) $x_{i,2} = I(i$ th person is from the West)

Qualitative Predictors with More than Two Levels

Using region to predict credit,

$$
y_i = \beta_0 + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \epsilon_i
$$

Why only two dummy variables? Consider:

- **E** If the *i*th person is from the South, $y_i = \beta_0 + \beta_1 x_{i,1} + \epsilon_i$.
- ▶ If the *i*th person is from the West, $y_i = \beta_0 + \beta_2 x_{i,2} + \epsilon_i$
- **►** If the *i*th person is from the East, $v_i = \beta_0 + \epsilon_i$

So each factor is represented in the model.

▶ Because East has no dummy variable, it is known as the baseline.

Qualitative Predictors with More than Two Levels

summary(**lm**(Limit **~** Region, data = credit))

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 4881.6 232.4 21.009 <2e-16 *** RegionSouth -153.1 284.3 -0.539 0.590 RegionWest -273.8 326.2 -0.839 0.402

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 Residual standard error: 2312 on 397 degrees of freedom Multiple R-squared: 0.001781, Adjusted R-squared: -0.003248 F-statistic: 0.3541 on 2 and 397 DF, p-value: 0.702

Qualitative Predictors

We can also use this approach for a mix of qualitative and quantitative variables in a model.

mod2 <- **lm**(Limit **~** Income **+** Rating **+** Own **+** Region, data=credit) **summary**(mod2)

Sometimes, two predictor variables interact in their impact on the outcome.

Example:

- ▶ Suppose spending money on TV advertising increases the effectiveness of radio advertising.
- $▶$ We want a way to let β_{radio} vary based on values of TV...

Accounting for Interactions

Consider

$$
Y=\beta_0+\beta_1X_1+\beta_2X_2+\beta_3X_1X_2+\epsilon
$$

How does this let β_{radio} vary based on values of $X_2 = TV$?

$$
Y = \beta_0 + (\beta_1 + \beta_3 X_2)X_1 + \beta_2 X_2 + \epsilon
$$

= $\beta_0 + \tilde{\beta}_1 X_1 + \beta_2 X_2 + \epsilon$

We can interpret *β*³ as the increase in effectiveness of TV advertising associated with a one-unit increase in radio advertising (or vice versa).

Consider: Why does estimating the coefficients not require any changes to our least squares approach?

Advertising <- **read.csv**("~/Courses/STAT 196M/datasets/Advertising.csv") mod3 <- **lm**(sales **~** TV **+** radio **+** TV*****radio, data=Advertising) **summary**(mod3)

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 6.750e+00 2.479e-01 27.233 <2e-16 *** TV 1.910e-02 1.504e-03 12.699 <2e-16 *** radio 2.886e-02 8.905e-03 3.241 0.0014 ** TV:radio 1.086e-03 5.242e-05 20.727 <2e-16 ***

Residual standard error: 0.9435 on 196 degrees of freedom Multiple R-squared: 0.9678, Adjusted R-squared: 0.9673 F-statistic: 1963 on 3 and 196 DF, p-value: < 2.2e-16

Consider R_{adj}^2 for the main effects model:

Advertising <- **read.csv**("~/Courses/STAT 196M/datasets/Advertising.csv") mod4 <- **lm**(sales **~** TV **+** radio, data=Advertising) **summary**(mod4)

Coefficients: Estimate Std. Error t value Pr(>|t|)

Residual standard error: 1.681 on 197 degrees of freedom Multiple R-squared: 0.8972, Adjusted R-squared: 0.8962 F-statistic: 859.6 on 2 and 197 DF, p-value: < 2.2e-16

In general, if we include an interaction term in a model, we also include the main effects even if the p-values associated with the main effects are not significant.

Consider Credit Balance predicted by Income and Student status.

▶ The interaction allows the model for students to have a different slope than the model for non-students, while the main effects model only allows for different intercepts.

Horsepower

Nonlinear Relationships Between Predictors and Outcome

How can we deal with this using linear regression?

- ▶ The model fit requires the model to be linear with respect to *β*.
- **•** This is much like including X_1X_2 in the model by creating a "new variable" in the matrix X.
- Here, we just construct a "new variable", say, X_1^2 in X.

Nonlinear Relationships Between Predictors and Outcome

mod5 <- **lm**(mpg **~ poly**(horsepower, 2), data = Auto) **summary**(mod5)

Coefficients: Estimate Std. Error t value Pr(>|t|) (Intercept) 23.4459 0.2209 106.13 <2e-16 *** poly(horsepower, 2)1 -120.1377 4.3739 -27.47 <2e-16 *** poly(horsepower, 2)2 44.0895 4.3739 10.08 <2e-16 *** Residual standard error: 4.374 on 389 degrees of freedom Multiple R-squared: 0.6876, Adjusted R-squared: 0.686 F-statistic: 428 on 2 and 389 DF, p-value: $\leq 2.2e-16$

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