3.2 Multiple Linear Regression

Prof. Lauren Perry

In practice, X is usually composed of more than one predictor variable.

- ▶ Multiple linear regression will allow us to deal with multiple inputs.
	- ▶ Want to put all useful inputs into the model at once.
- It also allows us to better model the case where the relationship between X and Y is not linear.

For p distinct predictors, the linear regression model takes the form

$$
Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p + \epsilon
$$

where X_j is the j th predictor and β_j quantifies the association between that variable and the response.

► We say that β_j is the average change in Y for a one unit increase in X_j , holding all other predictors fixed.

Estimating the Regression Coefficients

We make predictions using the formula

$$
\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots + \hat{\beta}_p x_p
$$

and we again estimate our parameters by minimizing the sum of squared residuals

$$
RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2
$$

the solutions to which are most easily represented using matrix algebra.

Estimating the Regression Coefficients

We will find these coefficients using R:

ads <- **read.csv**("~/Courses/STAT 196M/datasets/Advertising.csv") mod1 <- **lm**(sales **~** TV **+** radio **+** newspaper, data=ads) **round**(mod1**\$**coefficients,3)

(Intercept) TV radio newspaper ## 2.939 0.046 0.189 -0.001

(Recall that the advertising data is in thousands.)

Estimating the Regression Coefficients

summary(mod1)

produces the following:

Coefficient Estimate Std. Error t value Pr(>|t|) (Intercept) 2.938889 0.311908 9.422 <2e-16 *** TV 0.045765 0.001395 32.809 <2e-16 *** radio 0.188530 0.008611 21.893 <2e-16 *** newspaper -0.001037 0.005871 -0.177 0.86

Is at least one of the predictors X_1, X_2, \ldots, X_p useful in predicting the response Y?

- \blacktriangleright This is a little more complex than in the simple linear regression setting, where we could just examine *β*1.
- ▶ Here, we test

$$
H_0: \beta_1 = \beta_2 = \cdots = \beta_p = 0
$$

versus

 $H_{\sf a}$: at least one β_j is non-zero

This test uses an F-statistic:

$$
F = \frac{(TSS - RSS)/p}{RSS/(n-p-1)}
$$

where again TSS =
$$
\sum (y_i - \bar{y})^2
$$
 and RSS = $\sum (y_i - \hat{y}_i)^2$.

When there is no relationship between the predictors, we expect the F ratio to be close to 1.

In R, the command summary($mod1$) also produces the following: Residual standard error: 1.686 on 196 degrees of freedom Multiple R-squared: 0.8972, Adjusted R-squared: 0.8956 F-statistic: 570.3 on 3 and 196 DF, p-value: < 2.2e-16

Examining A Subset of Coefficients

Sometimes, we have reason to test whether a particular subset of q of the p coefficients are zero:

 H_0 : all of the q coefficients are zero

Here, we fit a second model that uses all of the variables *except* for the q variables of interest.

 \triangleright We call this model's residual sum of squares RSS₀. Then

$$
F = \frac{(\text{RSS}_0 - \text{RSS})/q}{\text{RSS}/(n-p-1)}
$$

- ▶ The p-values provided earlier in the coefficient output correspond to the setting where the single corresponding variable is omitted.
	- \triangleright i.e., the partial effect of adding that variable to the model.

If at least one coefficient has a small p-value, why do we still need to look at the overall F-statistic?

- \triangleright About 5% of the p-values associated with each variable will be below 0.05 just by chance.
- \triangleright So, with a lot of predictors, it's relatively likely that we would see small p-values even if there is no association between the predictors and the response.
	- \blacktriangleright The F-statistic adjusts for number of predictors, so it doesn't have this problem.
- \blacktriangleright Thus, we want to examine overall model fit as well as the significance of each coefficient.

Once we've decided the model is useful overall, we want to figure out which predictors are useful.

▶ We could just look at the p-values for each coefficient, but this can lead to some issues.

 \blacktriangleright Ex: if p is large, we may make some false discoveries.

 \blacktriangleright Instead, we use variable selection methods.

The ideal approach is to examine models for all possible subsets of the predictors.

We can then compare these models using statistics like

- 1. Mallow's C_p
- 2. Akaike information criterion (AIC)
- 3. Bayesian information criterion (BIC)
- 4. Adjusted R^2

These are studied more extensively in Chapter 6.

Unfortunately, examining all possible subsets isn't always feasible.

- \triangleright For $p = 2$ predictors, there are four possible models: $\blacktriangleright \space \boldsymbol{Y} = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$ $Y = \beta_0 + \beta_1x_1 + \epsilon$ $Y = \beta_0 + \beta_2 x_2 + \epsilon$ $Y = \beta_0 + \epsilon$
- \triangleright But the number of possible models grows quickly!
- \triangleright For p input variables, there are a total of 2^p possible subsets.

We need a way to automate variable selection that doesn't require us to examine all possible subsets.

Unfortunately, the methods discussed at this point in the textbook tend to lead to a variety of problems, so we will hold off on other options until Chapter 6.

Model Fit

- \blacktriangleright Correlation R and the coefficient of determination R^2 are conceptually the same for multiple linear regression.
- However, R^2 will *always* increase as more variables are added to the model.
- Instead, we will use an *adjusted* R^2 value that takes into account the number of input variables p.

$$
R_{adj}^2 = 1 - \left[\left(\frac{n-1}{n-p-1} \right) (1 - R^2) \right]
$$

- ▶ We interpret R_{adj}^2 the same way as R^2 .
- \triangleright This value is shown in the regression model summary output in R.

We can also examine graphical summaries for model fit.

Predictions

It's straightforward to plug in values of X to the estimated regression line.

Sources of error/uncertainty:

- 1. The coefficients are estimates, so $\hat{f}(X)$ is only an estimate for $f(X)$.
	- ▶ A source of reducible error.
	- \triangleright We can calculate confidence intervals for \hat{Y}
- 2. In practice, assuming linearity is probably only an approximation.
	- \triangleright Another source of reducible error, model bias.
	- \triangleright We generally ignore this if the model is "good enough".
- 3. Random error *ϵ*.
	- ▶ Irreducible error.
	- \triangleright We can also calculate *prediction intervals* for \hat{Y} .

Prediction Intervals

- \triangleright Confidence intervals quantify uncertainty for a *mean*.
	- ▶ For a 95% CI, we say that 95% of intervals of that form will contain the true value of $f(X)$.
	- \blacktriangleright I.e., the *average* outcome y for a point x.
- ▶ Prediction intervals quantify uncertainty for a *single point*.
	- ▶ For a 95% PI, we say that 95% of intervals of that form will contain the true value of Y for a specific point.
	- \blacktriangleright l.e., the *specific* outcome y for a point x.