# 3.2 Multiple Linear Regression

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In practice, X is usually composed of more than one predictor variable.

- Multiple linear regression will allow us to deal with multiple inputs.
  - Want to put all useful inputs into the model at once.
- It also allows us to better model the case where the relationship between X and Y is not linear.

For p distinct predictors, the linear regression model takes the form

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon$$

where  $X_j$  is the *j*th predictor and  $\beta_j$  quantifies the association between that variable and the response.

• We say that  $\beta_j$  is the average change in Y for a one unit increase in  $X_j$ , holding all other predictors fixed.

## Estimating the Regression Coefficients

We make predictions using the formula

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \dots + \hat{\beta}_p x_p$$

and we again estimate our parameters by minimizing the sum of squared residuals

$$\mathsf{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

the solutions to which are most easily represented using matrix algebra.

# Estimating the Regression Coefficients

```
We will find these coefficients using R:
```

```
ads <- read.csv("~/Courses/STAT 196M/datasets/Advertising.csv")
mod1 <- lm(sales ~ TV + radio + newspaper, data=ads)
round(mod1$coefficients,3)</pre>
```

## (Intercept) TV radio newspaper ## 2.939 0.046 0.189 -0.001

(Recall that the advertising data is in *thousands*.)

# Estimating the Regression Coefficients

summary(mod1)

produces the following:

Coefficient Estimate Std. Error t value Pr(>|t|) (Intercept) 2.938889 0.311908 9.422 <2e-16 \*\*\* ΤV 0.045765 0.001395 32.809 <2e-16 \*\*\* 0.188530 0.008611 21.893 <2e-16 \*\*\* radio 0.005871 0.86 newspaper -0.001037-0.177

Is at least one of the predictors  $X_1, X_2, \ldots, X_p$  useful in predicting the response Y?

- This is a little more complex than in the simple linear regression setting, where we could just examine β<sub>1</sub>.
- Here, we test

$$H_0:\beta_1=\beta_2=\cdots=\beta_p=0$$

versus

 $H_a$ : at least one  $\beta_j$  is non-zero

This test uses an F-statistic:

$$F = \frac{(\mathsf{TSS} - \mathsf{RSS})/p}{\mathsf{RSS}/(n - p - 1)}$$

where again 
$$TSS = \sum (y_i - \bar{y})^2$$
 and  $RSS = \sum (y_i - \hat{y}_i)^2$ .

When there is no relationship between the predictors, we expect the  ${\sf F}$  ratio to be close to 1.

In R, the command summary(mod1) also produces the following: Residual standard error: 1.686 on 196 degrees of freedom Multiple R-squared: 0.8972, Adjusted R-squared: 0.8956 F-statistic: 570.3 on 3 and 196 DF, p-value: < 2.2e-16</pre>

## Examining A Subset of Coefficients

Sometimes, we have reason to test whether a particular subset of q of the p coefficients are zero:

 $H_0$ : all of the q coefficients are zero

Here, we fit a second model that uses all of the variables except for the q variables of interest.

• We call this model's residual sum of squares  $RSS_0$ . Then

$$F = \frac{(\text{RSS}_0 - \text{RSS})/q}{\text{RSS}/(n - p - 1)}$$

- The p-values provided earlier in the coefficient output correspond to the setting where the single corresponding variable is omitted.
  - i.e., the partial effect of adding that variable to the model.

If at least one coefficient has a small p-value, why do we still need to look at the overall F-statistic?

- About 5% of the p-values associated with each variable will be below 0.05 just by chance.
- So, with a lot of predictors, it's relatively likely that we would see small p-values even if there is no association between the predictors and the response.
  - ▶ The F-statistic adjusts for number of predictors, so it doesn't have this problem.
- Thus, we want to examine overall model fit as well as the significance of each coefficient.

Once we've decided the model is useful overall, we want to figure out *which* predictors are useful.

We could just look at the p-values for each coefficient, but this can lead to some issues.

Ex: if p is large, we may make some false discoveries.

▶ Instead, we use *variable selection methods*.

The ideal approach is to examine models for all possible subsets of the predictors.

We can then compare these models using statistics like

- 1. Mallow's  $C_p$
- 2. Akaike information criterion (AIC)
- 3. Bayesian information criterion (BIC)
- 4. Adjusted  $R^2$

These are studied more extensively in Chapter 6.

Unfortunately, examining all possible subsets isn't always feasible.

- For p = 2 predictors, there are four possible models:  $Y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$   $Y = \beta_0 + \beta_1 x_1 + \epsilon$   $Y = \beta_0 + \beta_2 x_2 + \epsilon$  $Y = \beta_0 + \epsilon$
- But the number of possible models grows quickly!
- For p input variables, there are a total of  $2^p$  possible subsets.

We need a way to automate variable selection that doesn't require us to examine all possible subsets.

Unfortunately, the methods discussed at this point in the textbook tend to lead to a variety of problems, so we will hold off on other options until Chapter 6.

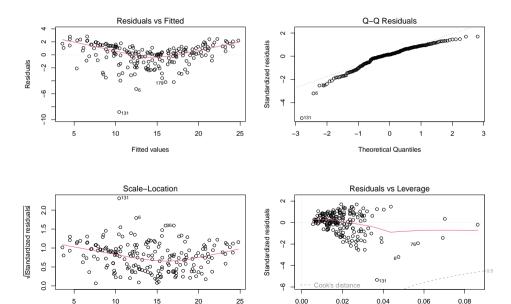
# Model Fit

- Correlation R and the coefficient of determination R<sup>2</sup> are conceptually the same for multiple linear regression.
- However,  $R^2$  will *always* increase as more variables are added to the model.
- Instead, we will use an *adjusted* R<sup>2</sup> value that takes into account the number of input variables p.

$$R_{adj}^2 = 1 - \left[ \left( rac{n-1}{n-
ho-1} 
ight) (1-R^2) 
ight]$$

- We interpret  $R_{adi}^2$  the same way as  $R^2$ .
- This value is shown in the regression model summary output in R.

We can also examine graphical summaries for model fit.



## Predictions

It's straightforward to plug in values of X to the estimated regression line.

Sources of error/uncertainty:

- 1. The coefficients are estimates, so  $\hat{f}(X)$  is only an estimate for f(X).
  - A source of reducible error.
  - We can calculate confidence intervals for  $\hat{Y}$ .
- 2. In practice, assuming linearity is probably only an approximation.
  - Another source of reducible error, *model bias*.
  - We generally ignore this if the model is "good enough".
- 3. Random error  $\epsilon$ .
  - Irreducible error.
  - We can also calculate *prediction intervals* for  $\hat{Y}$ .

## **Prediction Intervals**

- Confidence intervals quantify uncertainty for a mean.
  - For a 95% CI, we say that 95% of intervals of that form will contain the true value of f(X).
  - I.e., the average outcome y for a point x.
- Prediction intervals quantify uncertainty for a single point.
  - For a 95% PI, we say that 95% of intervals of that form will contain the true value of Y for a specific point.
  - I.e., the specific outcome y for a point x.